

1 **Weekday and weekend sedentary time and physical activity in differentially active**
2 **children**

3 ABSTRACT

4 Objectives. To investigate whether weekday-weekend differences in sedentary time (ST) and
5 specific intensities of physical activity exist among children categorised by physical activity
6 levels.

7 Design. Cross-sectional observational study.

8 Methods. Seven-day accelerometer data were obtained from 810 English children ($n = 420$
9 girls) aged 10-11 yr. Daily average $\text{min}\cdot\text{d}^{-1}$ spent in moderate to vigorous physical activity
10 (MVPA) were calculated for each child. Sex-specific MVPA quartile cut-off values
11 categorised boys and girls separately into four graded groups representing the least (Q1)
12 through to the most active (Q4) children. Sex- and activity quartile-specific multilevel linear
13 regression analyses analyzed differences in ST, light physical activity (LPA), moderate
14 physical activity (MPA), vigorous physical activity (VPA), and MVPA between weekdays
15 and weekends.

16 Results. On weekdays Q2 boys spent longer in LPA ($p < 0.05$), Q1 ($p < 0.001$), Q2 boys
17 ($p < 0.01$) did significantly more MPA, and Q1-Q3 boys accumulated significantly more VPA
18 and MVPA than at weekends. There were no significant differences in weekday and weekend
19 ST or physical activity for Q4 boys. On weekdays Q2 and Q3 girls accumulated more ST
20 ($p < 0.05$), Q1 and Q2 girls did significantly more MPA ($p < 0.05$), and Q1-Q3 girls engaged in
21 more VPA ($p < 0.05$) and more MVPA ($p < 0.01$) than at weekends. Q4 girls' ST and physical
22 activity varied little between weekdays and weekends.

23 Conclusions. The most active children maintained their ST and physical activity levels at
24 weekends, while among less active peers weekend ST and physical activity at all intensities
25 was lower. Low active children may benefit most from weekend intervention strategies.

26 Key words

27 High active, time-of-week differences, motor activity, health, child, multilevel analysis

28 Introduction.

29 Regular physical activity (PA) during childhood provides an array of health benefits ¹.

30 Insufficient childhood PA and excessive sedentary time (ST) however are independently

31 associated with negative health outcomes such as obesity and cardiometabolic risk ². To

32 establish healthy lifestyles in children efforts to increase PA and reduce ST are public health

33 priorities. To maintain good health, guidelines state that school-age youth accumulate at least

34 60 min•d⁻¹ of moderate-to-vigorous PA (MVPA) and reduce ST ¹.

35 Children's PA behaviours vary in bout frequency, duration, and intensity depending on the

36 setting where they occur. For example, there is greater uniformity in school day activity

37 compared to after-school and weekends, which allow more behavioural choice ³. Within such

38 contexts ST and PA levels can vary markedly due to the influence of multidimensional

39 correlates ⁴⁻⁵. Perhaps unsurprisingly, children tend to be less sedentary and active at

40 weekends than on weekdays ⁶⁻⁸. Weekends present more discretionary time for ST and PA,

41 but also lack the regular routines and structures of school weekdays which determine a

42 significant proportion of children's daily ST and PA ³. It is unclear though whether

43 engagement of high and low active children in ST and PA differs between weekdays and

44 weekends. Previous research found weekend PA to be lower than on weekdays regardless of

45 children's PA classification ⁹⁻¹¹, but seldom have such trends been examined by intensity-

46 specific PA ¹¹ or with ST as the outcome. Furthermore, weekends are an important context for

47 activity promotion but it is unknown whether targeting specific groups of children may be a

48 more efficient and efficacious intervention approach than a population-based one. Thus, the

49 study objectives were to investigate whether hypothesised differences in weekday and

50 weekend ST and specific intensities of PA exist among English children categorised by their

51 PA levels.

52 Methods.

53 Seventeen schools situated in a north-west England borough of over 300,000 people
54 participated in this cross-sectional research during 2008, 2009, and 2010. Eight schools were
55 recruited in 2008 and 2009 as part of a two year study of children's PA and health outcomes.
56 In 2010 three of these schools were joined by nine others for the baseline phase of a school-
57 based PA and nutrition intervention. In each year the same data collection procedures were
58 applied and so for the purposes of this paper the data were aggregated. To be eligible to
59 participate, children had to be in school Year 6 (aged 10-11 years) and be free from any
60 physical disabilities preventing them for taking part in routine physical activities. A fixed
61 available sample of all Year 6 children (N = 992; 307 in 2008, 295 in 2009, 390 in 2010)
62 were informed of the research by their class teachers, received project and consent
63 information, and were invited to participate. Written informed parental consent and child
64 assent were received from 818 children (230 in 2008, 270 in 2009, 318 in 2010; 82.5%
65 participation rate). Children participated at one time point only (i.e., 2008, 2008, or 2010).
66 The ethnic origin of the children was white British, which reflects the ethnic demographic of
67 the borough's school-age population ¹². In each year data were collected in one school per
68 week between October and December. Ethical approval was obtained from the University
69 Ethics Committee for each year of study (reference numbers 8.56 and 10/ECL/039). Ethical
70 principles of the Declaration of Helsinki were adhered to throughout this research.

71

72 Stature was measured to the nearest 0.1 cm using a portable stadiometer (Leicester Height
73 Measure, Seca, Birmingham, UK). Body mass was measured to the nearest 0.1 kg using
74 calibrated scales (Seca, Birmingham, UK) with the children in light clothing and barefooted.
75 Body mass index (BMI) was calculated (body mass (kg) / stature² (m²)) and BMI z-scores
76 were assigned to each child ¹³. International Obesity Task Force age and sex-specific BMI

77 cut-points were used to classify children as normal-weight or overweight/obese ¹⁴. Home
78 postal codes were used to generate indices of multiple deprivation (IMD) scores which
79 indicated area-level socio-economic status (SES). IMD scores are a composite of seven
80 domains of deprivation with higher scores representing higher degrees of deprivation ¹⁵.
81
82 PA was objectively measured using ActiGraph accelerometers (GT1M and GT3x, ActiGraph
83 LLC, Pensacola, FL) which were worn over the right hip for 7 consecutive days from waking
84 until bedtime. Accelerometers were set to record data using 5 second epochs. ActiGraph cut
85 points of 100 counts·min⁻¹, 101-2295 counts·min⁻¹, 2296 counts·min⁻¹, and 4012 counts·min⁻¹
86 classified the boundaries of ST, light intensity PA (LPA), moderate intensity PA (MPA) and
87 vigorous intensity PA (VPA), respectively ¹⁶. In the absence of universally agreed cut-points
88 to classify children's PA intensities, the cut-points of Evenson et al. ¹⁶ were selected on the
89 basis of a methodologically rigorous comparison study, which concluded that they have
90 acceptable classification accuracy across a range of intensities and are appropriate for use
91 with 5-15 year olds ¹⁷. MPA and VPA are influenced by different factors and both may affect
92 health outcomes differently ⁷. Taking these reasons together provided a rationale for studying
93 MPA and VPA separately from overall moderate to vigorous PA (MVPA). Non-
94 accelerometer wear time was defined as at least 20 min periods of consecutive zero counts ¹⁸.
95 Wear time criteria were at least 540 min·d⁻¹ on week days and 480 min·d⁻¹ on weekend days,
96 for at least two week days and one weekend day. These criteria have been shown to yield a
97 reliability of 0.9 suggesting a high degree of consistency across days ¹⁹. One hundred and
98 seventy seven children did not achieve the wear time criteria, and technical failures
99 downloading accelerometer data were experienced for a further 10 children. Missing data
100 analysis was completed using missing at random (MAR) assumptions. There is no way to
101 directly test these assumptions but relationships between health-related variables and missing

102 data can indicate whether they hold true ²⁰. In our sample there were no differences in BMI,
103 BMI z-score, weight status, and SES between the children who did and did not achieve the
104 accelerometer wear time criteria ($p>0.05$). On this basis we were satisfied that the data were
105 MAR rather than missing in a systematic manner. To replace the missing data, multiple
106 imputation ¹⁸ consisting of 100 iterations was undertaken separately for boys' and girls'
107 missing weekday and weekend accelerometer data values using the Markov Chain Monte
108 Carlo algorithm. Following five imputations, pooled values were generated and subsequently
109 integrated into the data set prior to analysis.

110

111 The daily average $\text{min}\cdot\text{d}^{-1}$ spent in MVPA was calculated for each child. Sex-specific MVPA
112 quartile cut-off values (Table 1) were calculated to categorise boys and girls separately into 4
113 graded groups representing the least through to the most active children. The least active
114 children were grouped in quartile (Q) 1 with the rest grouped in Q2, Q3, or Q4 (i.e., Q4 as the
115 most active group). Descriptive statistics were calculated for all measured variables and
116 preliminary comparisons were made between PA groups using ANOVAs, MANOVAs, and
117 χ^2 tests. The primary outcome variables were mean $\text{min}\cdot\text{d}^{-1}$ spent in ST, LPA, MPA, VPA,
118 and MVPA. These analyses and the multiple imputation were performed using IBM SPSS
119 Statistics version 20 (SPSS Inc., Chicago, IL) with alpha set at $p<0.05$.

120 For the main analyses multilevel linear regression analyses were computed separately for
121 boys and girls to analyze differences in outcome variables between weekdays and weekends.
122 Multilevel linear regression analyses accounted for the nested nature of the study design. A
123 three-level data structure was initially used where children were defined as the first level unit
124 of analysis, schools as the second level unit, and year of data collection as the third level unit
125 ²¹. Schools and years of data collection were included as levels in the analyses to control for
126 the effect that they could have on the children's activity ²¹. 'Crude' exploratory models were

127 initially constructed to determine whether schools and years of data collection were
128 significantly influential to warrant their inclusion as levels. If they were not, these levels were
129 removed resulting in models with either two or three level data structures. The necessity of
130 adding random slopes to these crude models was then evaluated by comparing differences
131 between the -2 log likelihood value with the value from the preceding model containing only
132 random intercepts, and computing a χ^2 test with two degrees of freedom ²¹. ‘Adjusted’
133 analyses were completed using the final crude models but which also controlled for the
134 effects of BMI z-score and SES, which are known to potentially confound differences in
135 children’s objectively measured PA ²². Data were analyzed using MLwiN 2.20 software
136 (Centre for Multi-Level Modelling, University of Bristol, UK). Regression coefficients in the
137 models were assessed for significance using the Wald statistic with one degree of freedom ²¹
138 and alpha was set at $p<0.05$.

139 Results.

140 Of the 818 children that returned written informed consent, three girls did not have BMI z-
141 scores due to missing anthropometric data, and three boys’ and two girls’ postcodes were
142 unavailable to calculate IMD scores. Thus, results were available from 810 children.
143 Participants were well matched in relation to age, stature, and IMD score regardless of PA
144 quartile (Table 1). Boys and girls in Q1 had significantly greater body mass than Q4 peers
145 ($p<0.05$), and a significantly higher percentage of Q1 boys were classed as overweight/obese,
146 than those in Q4 ($p<0.05$). Significantly more Q2-4 boys achieved recommended PA levels
147 than Q1 boys ($p<0.001$). Significantly fewer Q1-2 than Q3-4 girls met the PA guidelines
148 ($p<0.01$). Unadjusted weekday and weekend PA are presented in Table 2. As expected,
149 significant differences in ST, LPA, MPA, VPA, and MVPA were observed between Q1-4
150 boys, and these were generally most pronounced for weekdays ($p<0.05$ to 0.001). On

151 weekdays and weekends significant between quartile differences in girls ST, LPA, MPA,
152 VPA, and MVPA were also apparent ($p < 0.05$ to 0.001).

153 The main multilevel analyses investigated time of week differences in PA (Table 3). Q2 boys
154 spent longer in LPA on weekdays ($p < 0.05$). Q1 ($p < 0.001$) and Q2 boys ($p < 0.01$) did
155 significantly more weekday MPA than at weekends. Boys in Q1-Q3 accumulated
156 significantly more VPA and MVPA on weekdays than at weekends. In contrast, only small
157 differences in weekday and weekend ST and PA at all intensities were observed for Q4 boys.
158 Similar patterns emerged from analyses of the girls' data with the exception of ST, which Q2
159 and Q3 girls engaged in more on weekdays than at weekends ($p < 0.05$). Q1 and Q2 girls spent
160 significantly longer in weekday MPA than at weekends ($p < 0.05$). Girls in Q1-Q3 engaged in
161 more VPA ($p < 0.05$) and MVPA ($p < 0.01$) on weekdays than weekends. The exception to this
162 was Q4 girls who accumulated similar amounts of ST, and PA at all intensities between
163 weekdays and weekends.

164 Discussion.

165 This study found that the most active boys, and in particular, girls, maintained their ST and
166 PA levels across weekdays and weekends. In contrast Q1-Q3 children were more active and
167 less sedentary on weekdays than at weekends, with the exception of Q1-Q2 boys whose' ST
168 differed only by ~3 min. The decline in weekend PA observed in Q1-Q3 children concurs
169 with previous research demonstrating that early adolescents are less active and sedentary at
170 weekends regardless of age and gender ⁶. Moreover, for Q1-Q3 children the magnitude of
171 differences in weekday-weekend ST, LPA, MPA, VPA, and MVPA was generally similar,
172 suggesting that for most active children the weekend reductions in ST and PA were consistent
173 regardless of their PA level ⁹. Lower weekend ST and PA may be explained by various
174 factors. Sleep duration can limit waking hours available for recreational pursuits, and at

175 weekends this is longer than on weekdays²³. Further, the familial influence on ST and PA
176 (e.g., presence of siblings, social support, modelling) may be more significant at weekends
177 than at other periods of the week when discretionary opportunities for PA exist (e.g., after
178 school)^{5,22}. Parental concerns over child safety can also impact on weekend PA through
179 restrictions over children's local independent mobility²⁴.

180 A novel aspect of our study was comparing weekday and weekend ST and PA of children
181 classified by PA level. No previous studies have investigated ST in this way, and of the few
182 PA studies that have, the commonly observed trend of lower weekend PA was apparent.
183 Nyberg and colleagues compared weekday and weekend Actiwatch count·min⁻¹ between 6-10
184 year old Swedish children categorised into high, medium, and low activity tertiles⁹. Similar
185 declines in weekend PA for boys (14-19% decrease) and girls (13-24% decrease) were
186 reported regardless of activity tertile. In the ALSPAC study high and low active children
187 recorded marginally more ActiGraph count·min⁻¹ on weekdays than at weekends¹⁰. Neither
188 studies though reported PA defined by intensity. Using accelerometer count·min⁻¹ as a PA
189 outcome means that all recorded movement is defined as PA. This erroneously includes
190 episodes of ST as well as LPA. LPA in youth is inversely associated with cardiometabolic
191 health markers²⁵, although health gains from MVPA are greater²⁵, hence why this threshold
192 underpins PA recommendations for public health.

193 Intensity-specific PA was investigated by Jago et al. who observed a weekend reduction in
194 MVPA in the English children classed as 'high-active/low sedentary'¹¹. This contrasts with
195 our finding that the most active boys and girls either maintained or did not significantly
196 reduce their weekday MVPA at the weekend. The exact reasons for this are unclear and
197 previous literature on this specific topic is scarce. Favourable biological correlates such as
198 body composition, motor skill capability, and fitness are established factors that predispose
199 children to engage in PA and sport²⁶. Empirical evidence suggests that children and

200 adolescents involved in organised sport engage in more MVPA during weekdays and
201 weekends²⁷, and those who participate in competitive sport are more likely to meet PA
202 guidelines²⁸. Based on the higher VPA and MVPA levels of Q4 children it is plausible that a
203 significant proportion of their PA consisted of regular sport participation, some of which may
204 have been competitive. During the school week multiple opportunities for structured sports
205 and PA exist. When these opportunities are integrated into the school curriculum (e.g.,
206 physical education classes) and the extended school day (e.g., organised after-school
207 activities on school premises) they are accessible to all children regardless of sports
208 competence, thus optimising inclusivity and potential likelihood of participation²⁹. Out of
209 school weekday sporting opportunities however, are more likely be taken by children who
210 receive strong direct parental support (e.g., payment of club subscriptions, transport to and
211 from practice venues) and encouragement (e.g., positive verbal reinforcement and feedback)
212²², and demonstrate greater sport competence³⁰. Assuming Q4 children regularly engaged in
213 structured weekday sport that was partly facilitated by parental support and encouragement,
214 the known associations between sport participation and PA levels lead us to speculate that Q4
215 children may also have participated in weekend sports activities to a greater degree than their
216 less active peers²⁸. The role of parents may be a key factor in explaining these findings, as a
217 recent study reported how children with greater parental support had less of a decline in
218 weekend MPA and VPA over a one year period⁷. Though the role of parents and family is
219 consistent across studies it explains only a modest proportion of variance in weekend PA^{8,22}.
220 Thus, factors beyond the home and family are likely influences on children's weekend PA.
221 Future research should focus on establishing the multidimensional correlates of children's
222 weekend PA to better understand the determining factors of this behaviour.

223 Strengths of this study included the use of imputation procedures to maximise the sample
224 size. Furthermore, the high participation rate gave us confidence that the risk of participant

225 recruitment bias was minimized. Children were blinded to their classification into PA
226 quartiles which occurred retrospectively so as not to influence their activity behaviours. For
227 the same reason, researchers were blinded to the PA groupings during data collection. ST and
228 PA were measured objectively and the analyses accounted for potential confounding
229 variables and for the collection of data across different years in different schools. A study
230 limitation was the cross-sectional design which precluded us from knowing whether the
231 maintenance of the Q4 children's weekend ST and PA was consistent across other times of
232 the year and between years. Furthermore, the data were collected in autumn and winter when
233 children's ST is higher and PA is lower than in spring and summer⁵. For this reason
234 weekday-weekend patterns of ST and PA may differ when the climate is warmer and daylight
235 hours longer. The data were representative of predominantly white British 10-11 year olds in
236 one north-west England borough, and on this basis the findings cannot be generalised to other
237 groups or areas. PA was measured using hip-mounted accelerometry which has inherent
238 limitations related to accuracy of assessing PA on inclines, and during activities involving
239 cycling, swimming, and upper body movements. Further, lack of consensus as to the most
240 appropriate accelerometer PA cutpoints means that reported PA may have been
241 overestimated. Resource constraints though meant it was not possible to calculate
242 individually calibrated cutpoints which may provide more accurate PA data.

243 Conclusion.

244 ST and PA were highest on weekdays in all children except for the most active group who
245 maintained their ST and PA levels at weekends. These trends were particularly evident
246 among Q4 girls. Whether the most active children would continue to maintain their weekday-
247 weekend ST and PA patterns is unknown, and future research should investigate this through
248 prospective study designs which encompass multidimensional correlates of weekend ST and
249 PA. Such approaches may allow researchers to identify consistent correlates of weekend ST

250 and PA as well as specific periods of time when any declines in high active boys' and girls'
251 weekend ST and PA occur.

252 Practical implications.

- 253 • Weekends are important to target for interventions because observed decreases in
254 most children's PA are generally consistent regardless of age and gender.
- 255 • The least active children are most in need of weekend intervention strategies, through
256 cost-effective and resource efficient approaches.
- 257 • Parental support is important to attenuate reductions in weekend PA and to reduce
258 weekday ST.

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